

# LIMIT THEOREMS FOR SEMI-MARKOV PROCESSES WITH A COUNTABLE SET OF STATES

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**Abstract**

**Full Text**

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*MATHEMATICS*

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## LIMIT THEOREMS FOR SEMI-MARKOV PROCESSES WITH A COUNTABLE SET OF STATES

*(Presented by Academician V. M. Glushkov, 12 I 1970)*

Let, for each  $t \in (0, \infty)$ ,  $\nu_t(s) \in \{1, 2, \dots\}$  be a time-homogeneous right-continuous semi-Markov process (s.m.p.), which is specified, following <sup>(1, 2)</sup>, by the matrix of transition probabilities

$$F_t(i, j, u) = P\{\varepsilon_{k+1} = j, \tau_t(\varepsilon_k) < u \mid \varepsilon_k = i\}, \quad i, j = 1, 2, \dots,$$

where  $\varepsilon_k = \nu_t(\theta_t(k))$ ,  $\tau_t(\varepsilon_k) = \theta_t(k+1) - \theta_t(k)$ ,  $\theta_t(k)$  is the time of the  $k$ -th jump, i.e.  $\theta_t(0) = 0$ , and  $\theta_t(k) = \min\{s : s > \theta_t(k-1), \nu_t(s) \neq \varepsilon_{k-1}\}$ ,  $k \geq 1$ .

Introduce the following random variables:

$$\beta_t(i, k) = \min\{n : n > \beta_t(i, k-1), \varepsilon_n = i\}, \quad k \geq 1$$

$$(\beta_t(i, 0) = 0),$$

and  $\nu_t, \nu_t(i), \Omega_t(i), i = 1, 2, \dots$ , respectively, the total number of jumps of the s.m.p., the number of visits to state  $i$ , and the total time spent in  $i$  during time  $t$ .

Let, for each  $t \in (0, \infty)$ ,

$$f_t^{(k)}(i, x), \quad x \in (0, \infty), \quad i = 1, 2, \dots, \quad k = 0, 1, 2, \dots,$$

be a family of mutually independent random variables, independent of the s.m.p.  $\nu_t(s)$  (here and in what follows the quantities  $\gamma^{(k)}, k = 0, 1, 2, \dots$ , denote random variables independent and identically distributed with  $\gamma$ ). Introduce an additive functional  $S(t)$  of the form

$$S(t) = \sum_{k=0}^{\nu_t-1} f_t(\varepsilon_k),$$

where  $f_t(\varepsilon_k) = f_t^{(k)}(\varepsilon_k, \tau_t(\varepsilon_k))$ ,  $k = 0, 1, 2, \dots$

In what follows we shall assume that the embedded Markov chain, which, as is known, is specified by the matrix  $P(t) = \|p_t(i, j)\|$ ,  $i, j = 1, 2, \dots$ , where  $p_t(i, j) = F_t(i, j, \infty)$ ,  $i, j = 1, 2, \dots$ , has one positive class with stationary distribution  $q_t(i)$ ,  $i = 1, 2, \dots$ , and

$$M\tau_t(i) = m_t(i) < \infty, \quad i = 1, 2, \dots,$$

$$A_t = \sum_{i=1}^{\infty} q_t(i)m_t(i) < \infty,$$

and the matrix  $\bar{P} = \lim_{t \rightarrow \infty} P(t)$  also corresponds to a chain with one positive class.

Put

$$a_t = A_t^{-1} \sum_{i=1}^{\infty} q_t(i)m_t(f_i),$$

if  $Mf_t(i, \tau_t(i)) = m_t(f_i) < \infty$ ,  $i = 1, 2, \dots$ , and the given series converges absolutely, and  $a_t = 0$  otherwise. Denote

$$X_t(k) = \sum_{i=\beta_t(1,k)}^{\beta_t(1,k+1)-1} \tau_t(\varepsilon_i), \quad \varphi_t(k) = \sum_{i=\beta_t(1,k)}^{\beta_t(1,k+1)-1} f_t(\varepsilon_i),$$

$$Y_t(k) = \varphi_t(k) - a_t X_t(k), \quad k = 1, 2, \dots$$

**Theorem 1.** If there exist  $\gamma(t)$  and  $b(t)$  such that

$$P \left\{ \frac{1}{\gamma(t)} \sum_{k=1}^{V_t(1)} \left( X_t(k) - \frac{A_t}{q_t(1)} \right) < z \right\} \rightarrow F(z), * \quad (1)$$

$$P \left\{ \frac{1}{b(t)} \sum_{k=1}^{V_t(1)} Y_t(k) < z \right\} \rightarrow \Phi(z), \quad (2)$$

$$\gamma(t)/t \rightarrow 0, \quad A_t/\gamma(t)q_t(1) \rightarrow 0, \quad (3)$$

and for any  $\varepsilon > 0$ ,  $i = 1, 2, \dots$

$$P\{f_t(i, \tau_t(i)) > \varepsilon b(t)\} \rightarrow 0,$$

$$P\{a_t \tau_t(i) > \varepsilon b(t)\} \rightarrow 0$$

as  $t \rightarrow \infty$ , and  $F(z)$  is a proper distribution function, then, independently of the initial state of the s.m.p.,

$$P\left\{\frac{1}{b(t)}(S(t) - ta_t) < z\right\} \rightarrow \Phi(z).$$

**Corollary.** If  $a_t \neq 0$  and for any  $k = 1, 2, \dots$ , as  $\alpha_t = o(1)$ ,

$$M \exp\{i\lambda \alpha_t (m_t(f_k))^{-1} f_t(k, \tau_t(k))\} = 1 + \alpha_t i\lambda(1 + o_k(1)),$$

then, under assumptions (1) and (3), the quantity  $(ta_t)^{-1}S(t)$  tends in probability to one as  $t \rightarrow \infty$ , i.e., the law of large numbers is valid for  $S(t)$ .

Let us also note that under assumptions (1) and (3)

$$P\left\{\frac{A_t}{\gamma(t)q_t(1)}\left(v_t(1) - \frac{tq_t(1)}{A_t}\right) < z\right\} \rightarrow 1 - F(-z).$$

The significance of Theorem 1 is that it makes it possible to reduce the study of additive sums of an s.m.p. of the type  $S(t)$  to the study of sums of a nonrandom number of independent identically distributed summands.

We apply the results obtained to the study of the vector of numbers of visits and sojourn times in the states of a certain finite subset  $I = \{1, 2, \dots, r\}$  of the states of the s.m.p. To this end set

$$f_t(i, x) = \varphi_i + f_{ix}, \quad i = 1, \dots, r, \quad f_t(i, x) = 0, \quad i > r,$$

where  $\varphi_i, f_{ix}$ ,  $i = 1, \dots, r$ , are arbitrary real numbers.

Denote  $\theta_t^*(0) = 0$ , and

$$\theta_t^*(k) = \min\{\theta_t(l) : \theta_t(l) > \theta_t^*(k-1), \varepsilon_l \in I\}, \quad k \geq 1.$$

Set  $\tau_t^*(\varepsilon_k^*) = \theta_t^*(k+1) - \theta_t^*(k)$ , where  $\varepsilon_k^* = \chi_t(\theta_t^*(k))$ . It is clear that

$$\tau_t^*(i) = \tau_t(i) + \tilde{\tau}_t(i), \quad i \in I,$$

where

$$\tilde{\tau}_t(i) = \min\{s : \chi_t(s) \in I \mid \chi_t(0-0) = i, \chi_t(0) \neq i\}.$$

\* Here the notation  $V_t(j)$  denotes the integer part of  $\frac{t}{A_t}q_t(j)$ ,  $j = 1, 2, \dots$

Suppose that for each  $j \in I$  there exist  $b_t(j)$  and  $B_t(j)$  such that the joint distribution

$$\left( \frac{1}{b_t(j)} \sum_{k=1}^{\nu_t(j)} (\tau_t^{(k)}(j) - m_t(j)), \frac{1}{B_t(j)} \sum_{k=1}^{\nu_t(j)} (\tilde{\tau}_t^{(k)}(j) - \tilde{m}_t(j)) \right) \xrightarrow{\text{sl}} (\xi_j, \tilde{\xi}_j)^*, \quad (4)$$

$$D_t/t \rightarrow 0, \quad A_t/D_t \rightarrow 0,$$

where

$$D_t = \max\{\sqrt{tA_t}, b_t(j), B_t(j), j \in I\},$$

$$m_t = \max\{m_t(j), j \in I\}, \quad \tilde{m}_t(j) = M\tilde{\tau}(j).$$

Let

$$\psi_j(\lambda_1, \lambda_2) = M \exp\{i(\lambda_1 \xi_j + \lambda_2 \tilde{\xi}_j)\}.$$

**Theorem 2.** If condition (4) is fulfilled and  $tm_t(A_t b_t)^{-1} \rightarrow 0$ , then the joint distribution of the random vector

$$\left\{ \frac{m_t}{B_t} \left( \nu_t(1) - \frac{tq_t(1)}{A_t} \right), \sqrt{\frac{A_t}{t}} \left( \nu_t(1) - \frac{q_t(1)}{q_t(i)} \nu_t(i) \right), \right. \\ \left. i = 2, \dots, r, g_t(k) \left( \Omega_t(k) - \frac{tq_t(k)m_t(k)}{A_t} \right), k = 1, \dots, r \right\},$$

independently of the initial state, converges weakly to a distribution with characteristic function of the form

$$\exp \left\{ -\frac{\sigma^2(\varphi_2, \dots, \varphi_r)}{2} \right\} \prod_{j=1}^r \psi_j(\rho_j f_j - \tilde{\rho}_j a(f), \alpha_j a(f)),$$

where the parameters  $\varphi_i(f_i)$ ,  $i = 1, \dots, r$ , correspond to the first (last)  $r$  components of the vector,  $\sigma^2(\varphi_2, \dots, \varphi_r)$  is a nondegenerate quadratic form in the variables  $\varphi_i$ ,  $i = 2, \dots, r$ . Here

$$B_t = \max \left\{ b_t(i), \frac{m_t}{A_t} B_t(i), i = 1, \dots, r \right\}, \quad a(f) = \varphi_1 + \sum_{i=1}^r c_i f_i,$$

$g_t(k)$ ,  $k = 1, \dots, r$ , are normalizing factors, and  $\rho_j, \tilde{\rho}_j, \alpha_j, c_i$ ,  $j = 1, \dots, r$ , are certain constants.

The results obtained, in particular, as applied to countable Markov chains having an ergodic distribution, agree with the results of V. A. Volkonskii<sup>(3)</sup>, and for a p.m.p. for which

$$F_t(i, j, u) = p(i, j)F(i, u), \quad i, j = 1, 2, \dots,$$

under our assumptions they make it possible to obtain the corresponding results of H. Kesten<sup>(4)</sup>.

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\* In the sense of weak convergence of distribution functions.

*Note: Figure translations are in progress. See original paper for figures.*

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