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DISTRIBUTIONS OF  
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SEMI-MARKOV  
PROCESS DEFINED ON  
A FIXED SET OF  
STATES UP TO THE  
MOMENT OF FIRST  
EXIT**

MATHEMATICS

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**Abstract**

**Full Text**

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*MATHEMATICS*

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## LIMITING DISTRIBUTIONS OF FUNCTIONALS OF A SEMI-MARKOV PROCESS DEFINED ON A FIXED SET OF STATES UP TO THE MOMENT OF FIRST EXIT

*(Presented by Academician V. M. Glushkov, January 12, 1970)*

Consider, for each  $t \in (0, \infty)$ , a right-continuous semi-Markov process (S.M.P.)  $\varkappa_t(s) \in \{0, 1, \dots, r\}$ , which is specified, following <sup>(1,2)</sup>, by the matrix of transition probabilities

$$F_t(i, j, u) = P\{\varepsilon_{k+1} = j, \tau_t(\varepsilon_k) < u \mid \varepsilon_k = i\}, \quad i, j = 0, \dots, r,$$

where  $\varepsilon_k = \varkappa_t(\theta_t(k))$ ,  $\tau_t(\varepsilon_k) = \theta_t(k+1) - \theta_t(k)$ , and  $\theta_t(k)$  is the moment of the  $k$ -th jump, i.e.  $\theta_t(0) = 0$ , and

$$\theta_t(k) = \min\{s : s > \theta_t(k-1), \varkappa_t(s) \neq \varepsilon_{k-1}\}, \quad k \geq 1.$$

We shall call the set of states  $\langle \omega \rangle = \langle 1, 2, \dots, r \rangle$  a group if, for any  $i, j \in \langle \omega \rangle$ ,  $p_t(i, j, \langle \omega \rangle) \rightarrow 1$  as  $t \rightarrow \infty$ , where  $p_t(i, j, \langle \omega \rangle)$  is the probability of reaching  $j$  from  $i$  before the S.M.P. leaves  $\langle \omega \rangle$ .

Let, for each  $t \in (0, \infty)$ ,  $f_t^{(k)}(i, x)$ ,  $k = 0, 1, \dots$ ,  $i \in \langle \omega \rangle$ ,  $x \in (0, \infty)$ , be a family of random variables independent of the S.M.P.  $\varkappa_t(s)$ , independent in the aggregate, whose distribution does not depend on the index  $k$ . Denote

$$\nu_t(j) = \min\{n : \varepsilon_n \in \overline{\langle \omega \rangle} \mid \varepsilon_0 = j\}, \quad j \in \langle \omega \rangle.$$

Introduce an additive functional of the trajectory of the S.M.P. of the form

$$\xi_t(j) = \sum_{k=0}^{\nu_t(j)-1} f_t^{(k)}(\varepsilon_k, \tau_t(\varepsilon_k)), \quad \text{if } \varepsilon_0 = j.$$

We shall study the limiting distributions of  $\xi_t(j)$  under the corresponding normalization, as  $t \rightarrow \infty$ . Denote by  $\tau_t(l, j)$  a random variable with distribution function  $p_t(l, j)^{-1} F_t(l, j, u)$ ,  $l, j \in \langle \omega \rangle$ , and let

$$\gamma_t(l, j) = f_t^{(1)}(l, \tau_t(l, j)), \quad \psi_t(l, j, \lambda) = M \exp\{i\lambda\gamma_t(l, j)\}.$$

Put

$$g_t = \sum_{k \in \langle \omega \rangle} q_t(k, \langle \omega \rangle) \sum_{i \in \langle \omega \rangle} p_t(k, i),$$

where  $q_t(k, \langle \omega \rangle)$ ,  $k \in \langle \omega \rangle$ , is the stationary distribution for the chain with matrix  $\tilde{P}(t, \langle \omega \rangle) = \|\tilde{p}_t(k, j, \langle \omega \rangle)\|$ ,  $k, j \in \langle \omega \rangle$ , and

$$\tilde{p}_t(k, j, \langle \omega \rangle) = p_t(k, j) \left( 1 - \sum_{i \in \langle \omega \rangle} p_t(k, i) \right)^{-1}.$$

Here  $p_t(k, j) = F_t(k, j, \infty)$ ,  $k, j = 0, \dots, r$ , are the transition probabilities for the embedded Markov chain. Suppose that there exists  $\beta = \beta(g_t) \rightarrow 0$  such that, for any  $l, j \in \langle \omega \rangle$ ,

$$\psi_t(l, j, \beta\lambda) = 1 + a_{ij}(\lambda) g_t(q_t(l, \langle \omega \rangle) p_t(l, j))^{-1} (1 + o(1)), \quad (1)$$

where  $|a_{ij}(\lambda)| \geq 0$ . Note that this condition is equivalent to the condition that the quantities  $\gamma_t(l, j)$  are attracted to certain infinitely divisible laws. Denote  $\varphi_t(j, \lambda) = M \exp\{i\lambda\xi_t(j)\}$ .

**Theorem 1.** *If the set  $\langle \omega \rangle$  forms a group and (1) is satisfied, then, independently of the initial state  $j \in \langle \omega \rangle$ , as  $t \rightarrow \infty$*

$$\varphi_t(j, \beta\lambda) \rightarrow \frac{1}{1 - a(\lambda)};$$

here

$$a(\lambda) = \sum_{i, j \in \langle \omega \rangle} a_{ij}(\lambda)$$

is the logarithm of the characteristic function of a certain infinitely divisible law.

**Remark 1.** In the case when  $a_{ij}(\lambda) = c_{ij}\lambda^\alpha$ ,  $i, j \in \langle \omega \rangle$  ( $c_{ij}$  are certain constants), the theorem has a simple probabilistic meaning:

$$\beta \xi_t(j) \xrightarrow{\text{sl}} \xi \eta^{1/\alpha *},$$

where  $\xi$  is a stable distribution with exponent  $\alpha$ ,  $0 < \alpha \leq 2$ ,  $\eta$  is an exponential distribution, and  $\xi$  and  $\eta$  are independent.

**Remark 2.** If, for the quantities  $\gamma_t(k, j)M_t^{-1}$ ,  $k, j \in \langle \omega \rangle$ , condition (1) is satisfied, where  $a_{kj}(\lambda) = i\lambda c_{kj}$ , and, in addition, the quantities  $\gamma_t(k, j)$  have an  $n$ -th moment ( $n \geq 1$ ), with

$$M_t^{-l} g_t^{l-1} q_t(k, \langle \omega \rangle) p_t(k, j) M \gamma_t(k, j)^l = o(1),$$

$$1 < l \leq n, \quad k, j \in \langle \omega \rangle,$$

then under the conditions of Theorem 1

$$M_t^{-1} g_t \xi_t(j) \xrightarrow{\text{sl}} \eta$$

and all moments up to order  $n$ , inclusive, converge. Here

$$M_t = \sum_{i \in \langle \omega \rangle} q_t(i, \langle \omega \rangle) M f_t^{(1)}(i, \tau_t(i)),$$

$\eta$  is exponential and  $M\eta = 1$ . In particular, if we set  $f_t^{(k)}(i, x) = 1$  or  $f_t^{(k)}(i, x) = x$ ,  $i \in \langle \omega \rangle$ , we obtain that

$$g_t \nu_t(j) \xrightarrow{\text{sl}} \eta, \quad m_t^{-1} g_t \Omega_t(i) \xrightarrow{\text{sl}} \eta,$$

where  $\Omega_t(j)$  is the time spent by the s.m.p. in the group  $\langle \omega \rangle$  before exit, under the condition that  $\varepsilon_0 = j \in \langle \omega \rangle$ , and

$$m_t = \sum_{i \in \langle \omega \rangle} q_t(i, \langle \omega \rangle) M \tau_t(i).$$

This agrees with the results of (3, 4).

Let us now assume that the quantities  $\gamma_t(k, j)$  have a first moment and that there exists  $\beta = \beta(g_t)$  such that, for all  $k, j \in \langle \omega \rangle$ ,

$$\psi_t(k, j, \beta \lambda) = 1 + i\lambda m_t(k, j)\beta + a_{kj}(\lambda) g_t \frac{1 + o(1)}{q_t(k, \langle \omega \rangle) p_t(k, j)}. \quad (2)$$

**Theorem 2.** *If the set  $\langle \omega \rangle$  forms a group and (2) is satisfied, then, independently of the initial state  $l \in \langle \omega \rangle$ , the distribution of the random variable*

$$\alpha_t(\xi_t(l) - M_t \nu_t(l))$$

*converges weakly to the distribution with characteristic function*

$$\frac{1}{1 - c_1 a(\lambda) + c_2 \lambda^2},$$

\* In the sense of weak convergence of distribution functions.

where  $\alpha_t = \min\{\sqrt{g_t} M_t^{-1}, \beta\}$ ;  $a(\lambda)$  is defined as in Theorem 1, and  $c_1 \geq 0$  and  $c_2 \geq 0$  are certain constants, with, if

$$\beta^2 M_t^2 g_t^{-1} \rightarrow \mu,$$

then for  $\mu = \infty$ ,  $c_1 = 0$ , and for  $\mu = 0$ ,  $c_2 = 0$ .

In conclusion we indicate an effective algorithm which, in a finite number of steps, not exceeding  $r - 1$ , makes it possible to check whether the given set  $\langle \omega \rangle$  will form a group.

Call a transition from  $i$  to  $j$  “proper” if  $P_t(i, j) \neq 0$ . We perform the following operation with our embedded chain: at the first step we carry out all “proper” transitions. We obtain a certain directed graph. It is obvious that connected subsets from which there are no transitions form groups. Replace each such group  $\langle d \rangle$  by one state with transition probabilities

$$p_t(\langle d \rangle, j) = \left( \sum_{k \in \langle d \rangle} q_t(k, \langle d \rangle) \sum_{l \in \langle d \rangle} p_t(k, l) \right)^{-1} \sum_{i \in \langle d \rangle} q_t(i, \langle d \rangle) p_t(i, j);$$

$$p_t(j, \langle d \rangle) = \sum_{k \in \langle d \rangle} p_t(j, k); \quad j \in \overline{\langle d \rangle}.$$

Thus, after the first step we obtain a new chain with a smaller number of states. With it we perform the same operation. Then it is clear that  $\langle \omega \rangle$  forms a group if and only if all of  $\langle \omega \rangle$  can be reduced to a single state.

It is of interest to note that the types of distributions found were indicated earlier in <sup>(5, 6)</sup> and, through the construction of  $\xi_t(j)$ , are naturally connected with the problem considered there.

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*Note: Figure translations are in progress. See original paper for figures.*

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