

# ON THE NUMBER OF ENTRIES OF TRAJECTORIES OF A MARKOV PROCESS INTO A GIVEN SET OF THE PHASE SPACE

MATHEMATICS

1969

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**Abstract**

**Full Text**

UDC 519.217

*MATHEMATICS*

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## ON THE NUMBER OF ENTRIES OF TRAJECTORIES OF A MARKOV PROCESS INTO A GIVEN SET OF THE PHASE SPACE

*(Presented by Academician A. N. Kolmogorov, 20 XII 1968)*

Let  $\xi_t$ ,  $-\infty < t < +\infty$ , be a stationary, time-homogeneous Markov process with values in the measurable space  $[\mathcal{E}, \mathfrak{B}_{\mathcal{E}}]$ .

**Definition 1.** We shall say that the time  $t$  is an **entry of a trajectory** of the process  $\xi_t$  into  $B$  from  $\bar{A}$ , if  $\xi_t \in B \subseteq A$  and, for any  $\varepsilon > 0$ , there exist  $s \in (t - \varepsilon, t)$  such that  $\xi_s \in \bar{A} = \mathcal{E} \setminus A$ .

By  $\eta_A(\Delta, B)$  we denote the number of entries of  $\xi_t$  into  $B$  from  $\bar{A}$  that occurred in the time interval  $\Delta$ , assuming that  $\eta_A(\Delta, B)$  are random variables. The random variables  $\eta_A(\Delta, B)$  form a stationary random stream <sup>(1)</sup>, whose characteristics (intensity, moments, etc.) can be expressed in terms of the function  $M_{x,A}(\Delta, B)$ , defined for  $\Delta \subset R^+ = (0, +\infty)$  by the relation

$$M_{x,A}(\Delta, B) = M\{\eta_A(\Delta, B) \mid \xi_0 = x\}. \quad (1)$$

Below we shall call this function **basic**. Suppose that  $M_{x,A}(\Delta, A) < \infty$  for every interval  $\Delta \subset R^+$  having finite length  $|\Delta| < \infty$ . This assumption is essential; however, it is satisfied for a wide class of Markov processes with a discrete intervention of chance <sup>(2)</sup> and sets  $B \subseteq A \subset \mathcal{E}$ . Under the assumption made,  $M_{x,A}(\Delta, B)$  can be extended to a measure on  $[A \times R^+, \mathfrak{B}_A \times \mathfrak{B}_{R^+}]$ , where  $\mathfrak{B}_A = \mathfrak{B}(B : B \subseteq A, B \in \mathfrak{B}_{\mathcal{E}})$ ,  $\mathfrak{B}_{R^+}$  is the  $\sigma$ -algebra of Borel sets on  $R^+$ . The basic functions satisfy the relation

$$M_{x,A}(\Delta + t, B) = \int_{y \in \mathcal{E}} P_x(t, dy) M_{y,A}(\Delta, B), \quad (2)$$

where  $\Delta + t = \{s : s - t \in \Delta\}$ ,  $\Delta \subset R^+$ , and  $P_x(t, A)$  are the transition probability functions.

**Definition 2.** A family of measures  $P_t$  on  $[\mathcal{E}, \mathfrak{B}_{\mathcal{E}}]$  **converges**, as  $t \rightarrow \infty$ , **to  $P$  with respect to the partition  $\mathcal{C} = \{\mathcal{E}_{n,k}\}$** ,  $\mathcal{E} = \bigcup_k \mathcal{E}_{n,k}$ ,  $\mathcal{E}_{n,k} \cap \mathcal{E}_{n,l} = \emptyset$ ,  $k \neq l$ ,  $\mathcal{E}_{n,k} = \bigcup_{l \in I_{n,k}} \mathcal{E}_{n+1,l}$ , if for every  $n$

$$\lim_{t \rightarrow \infty} \sum_k |P_t(\mathcal{E}_{n,k}) - P(\mathcal{E}_{n,k})| = 0. \quad (3)$$

From the convergence of  $P_t$  to  $P$  in variation (3) there follows convergence with respect to any partition  $\mathcal{C}$ .

**Theorem 1.** Let  $P_x(t, A)$  converge to the stationary measure  $P_{\infty}(A) = P\{\xi_t \in A\}$  with respect to  $\mathcal{C}$ , and let the basic functions be such that

$$\lim_{n \rightarrow \infty} \max_k \left[ \sup_{x \in \mathcal{E}_{n,k}} M_{x,A}(\Delta_0 + t_0, B) - \inf_{x \in \mathcal{E}_{n,k}} M_{x,A}(\Delta_0 + t_0, B) \right] = 0, \quad (4)$$

$$M_{x,A}(\Delta_0 + t_0, B) \leq g(x),$$

$$\lim_{t \rightarrow \infty} \int_{y \in \mathcal{E}} P_t(x, dy) g(y) = \int_{\mathcal{E}} P_{\infty}(dy) g(y).$$

Then

$$\lim_{t \rightarrow \infty} M_{x,A}(\Delta_0 + t, B) = \mu_A(B) = M\eta_A(\Delta_0, B), \quad (5)$$

where  $\Delta_0 = (0, 1)$ .

**Proof** follows from the uniform integrability of  $M_{y,A}(\Delta_0 + t_0, B)$ , the approximation of the main function by step functions on  $\mathcal{C}$ , and relations (2), (3).

Relation (5) is an analogue of Blackwell's theorem (4). The main difficulty in using Theorem 1 consists in verifying the convergence of  $P_t(x, A)$  to  $P_{\infty}(A)$  with respect to partitions  $\mathcal{C} = \{\mathcal{C}_{n,k}\}$  that are "sufficiently" fine for (4).

If  $\Delta_i = (t_i, t_i + u_i)$ ,  $t_i + u_i < t_{i+1}$  ( $\Delta_i < \Delta_{i+1}$ ),  $i = 1, \dots, k$ , then to the  $k$ -dimensional rectangle  $\Delta_1 \times \dots \times \Delta_k$  we assign the number

$$\begin{aligned} & M_{(k),A}(\Delta_1 \times \dots \times \Delta_k, B) = \\ &= \int_{y_1 \in B} \int_{s_1 \in \Delta_1} \mu_A(dy_1) ds_1 \int_{y_2 \in B} \int_{s_2 \in \Delta_2 - s_1} M_{y_1,A}(ds_2, dy_2) \dots \\ & \dots \int_{y_k \in B} \int_{s_k \in \Delta_k - (s_1 + \dots + s_{k-1})} M_{y_{k-1},A}(ds_k, dy_k). \end{aligned}$$

We shall assume that  $M_{(k),A}$  is bounded for any bounded intervals  $\Delta_i < \Delta_{i+1}$ ,  $i = 1, \dots, k$ . In this case it can be extended to a measure on the Borel sets of the cone  $C_k = \{(t_1, \dots, t_k), t_1 < t_2 < \dots < t_k\}$ . Using permutations of the values of the coordinates of the points of  $C_k$ , one can extend  $M_{(k),A}$  to a measure in  $R^k$ . The introduced measure has the property that, for any Borel set  $D \subset R^k$ , the measure  $M_{(k),A}(D, B)$  is equal to the mathematical expectation of the number of all points  $(t_1, \dots, t_k) \in D$  whose coordinates  $t_i$ ,  $i = 1, \dots, k$ , are entrances of  $\xi_t$  into  $B$  from  $\bar{A}$ .

**Definition 3.** The measure  $M_{(k),A}$  will be called the  $k$ -leading measure of the flow of entrances into  $B$  from  $\bar{A}$ .

Integrating the  $k$ -leading measure over appropriate sets in  $R^k$ , one can obtain various moments of the random variables  $\eta_A(\Delta, B)$  (see (5)). For example, the following holds.

**Theorem 2.** *The  $k$ -th factorial moment of the number of entrances into  $B$  from  $\bar{A}$  during the time  $(0, t)$  is equal to*

$$M_{(k)}(t, B, A) = k! \int_{y_1 \in B} \int_0^t \mu_A(dy_1) ds_1 \cdots \int_{y_k \in B} \int_0^{t-s_1-\dots-s_{k-1}} M_{y_{k-1}, A}(ds_k, dy_k). \quad (6)$$

**Proof** follows from the fact that the number of selections  $(t_1, \dots, t_k)$ ,  $t_1 < t_2 < \dots < t_k$ , from  $\eta_A((0, t), B) = \eta$  entrances into  $B$  from  $\bar{A}$  is equal to  $\eta(\eta-1) \cdots (\eta-k+1)/k!$

Studying the asymptotics of integrals of  $k$ -leading measures of flows of entrances into  $B$  from  $\bar{A}$ , one can obtain limit theorems on the convergence of “thinned” flows to Poisson flows.

**Definition 4.** We shall say that a random flow  $\eta_\varepsilon(\Delta)$ ,  $\varepsilon > 0$ , converges in distribution to a flow  $\eta_0(\Delta)$  as  $\varepsilon \rightarrow 0$  ( $\eta_\varepsilon(\Delta) \Rightarrow \eta_0(\Delta)$ ), if for any collection of intervals  $\Delta_1, \dots, \Delta_m$ ,

$$\lim_{\varepsilon \rightarrow 0} P_\varepsilon\{\eta_\varepsilon(\Delta_i) = k_i, i = 1, \dots, m\} = P_0\{\eta_0(\Delta_i) = k_i, i = 1, \dots, m\}.$$

Let  $\xi_{t,\varepsilon}$  be a stationary Markov process on  $[\mathcal{E}_\varepsilon, \mathfrak{B}_\varepsilon]$ , and let  $\eta_{A_\varepsilon}(\Delta, B_\varepsilon)$  be the flow of entrances into  $B_\varepsilon$  from  $\bar{A}_\varepsilon$ ,  $\varepsilon > 0$  a parameter, and  $M_{x, A_\varepsilon}(\Delta, B_\varepsilon)$  the main function of the flow  $\eta_{A_\varepsilon}(\Delta, B_\varepsilon)$ . If  $\Delta = (t_1, t_2)$ , then put

$$\Delta/\mu_{A_\varepsilon}(B_\varepsilon) = (t_1/\mu_{A_\varepsilon}(B_\varepsilon), t_2/\mu_{A_\varepsilon}(B_\varepsilon)).$$

Introduce the majorant

$$\overline{M}_{A_\varepsilon}(B_\varepsilon) = \sup_{s>0} \sup_{y \in B_\varepsilon} M_{y, A_\varepsilon}(\Delta_0 + s, B_\varepsilon). \quad (7)$$

**Theorem 3.** If the fundamental functions  $M_{y, A_\varepsilon}(\Delta, B_\varepsilon)$  satisfy the conditions

$$\lim_{\varepsilon \rightarrow 0} \overline{M}_{A_\varepsilon}(B_\varepsilon) = 0$$

and, uniformly in  $y \in B_\varepsilon$  and  $\varepsilon \leq \varepsilon_0$ ,

$$\lim_{s \rightarrow \infty} M_{y, A_\varepsilon}(\Delta_0 + s, B_\varepsilon) / \mu_{A_\varepsilon}(B_\varepsilon) = 1, \quad (9)$$

then

$$\eta_{A_\varepsilon}(\Delta / \mu_{A_\varepsilon}(B_\varepsilon), B_\varepsilon) \Rightarrow \eta_0(\Delta), \quad (10)$$

where  $\eta_0(\Delta)$  is a Poisson flow with unit intensity.

The **proof** is based on verifying the convergence of the moments of the random variables  $\eta_{A_\varepsilon}(\Delta_i / \mu_{A_\varepsilon}(B_\varepsilon), B_\varepsilon)$ ,  $\Delta_i \cap \Delta_j = \emptyset$ ,  $i \neq j$ ,  $i, j = 1, \dots, k$ ,  $k = 1, 2, \dots$ , to the moments of mutually independent Poisson random variables with means equal to  $|\Delta_i|$ . For example, for  $k = 1$  it is first proved that asymptotically the integral (6), with  $A, B, t$  replaced by  $A_\varepsilon, B_\varepsilon, t_\varepsilon = t / \mu_{A_\varepsilon}(B_\varepsilon)$ , is equivalent to the integral

$$\begin{aligned} M_{(k)}(\tau_\varepsilon, t_\varepsilon, B_\varepsilon, A_\varepsilon) &= k! \int_{y_1 \in B_\varepsilon} \int_0^{t_\varepsilon - (k-1)\tau_\varepsilon} \mu_{A_\varepsilon}(dy_1) ds_1 \times \\ &\times \int_{y_2 \in B_\varepsilon} \int_{\tau_\varepsilon}^{t_\varepsilon - s_1 - (k-2)\tau_\varepsilon} M_{y_1, A_\varepsilon}(ds_2, dy_2) \cdots \int_{y_k \in B_\varepsilon} \int_{\tau_\varepsilon}^{t - s_1 - \cdots - s_{k-1}} M_{y_{k-1}, A_\varepsilon}(ds_k, dy_k) \end{aligned}$$

for a corresponding growth of  $\tau_\varepsilon$ . From (8) we find that, uniformly in  $y \in B_\varepsilon$ ,

$$\lim_{\varepsilon \rightarrow 0, \tau_\varepsilon \rightarrow \infty, u_\varepsilon \rightarrow \infty} \int_{\tau_\varepsilon}^{\tau_\varepsilon + u_\varepsilon} (\tau_\varepsilon + u_\varepsilon - s)^k M_{y, A_\varepsilon}(ds, B_\varepsilon) / \mu_{A_\varepsilon}(B_\varepsilon) (u_\varepsilon^{k+1} / k + 1) = 1.$$

Hence, in turn, we obtain that

$$\lim_{\varepsilon \rightarrow 0} M(\tau_\varepsilon, t_\varepsilon, B_\varepsilon, A_\varepsilon) = t^k.$$

The convergence of the finite-dimensional distributions of the quantities  $\eta_{A_\varepsilon}(\Delta_i/\mu_{A_\varepsilon}(B_\varepsilon), B_\varepsilon)$  to Poisson distributions is obtained as a consequence of the theorem on convergence of moments [6].

**Definition 5.** An entrance  $t$  into  $B$  from  $\bar{A}$  will be called a  $\tau$ -entrance ( $\bar{\tau}$ -entrance) if in the interval  $(t - \tau, t)$  there are no entrances (there are entrances) into  $B$  from  $\bar{A}$ . A  $\tau$ -entrance into  $B$  from  $\bar{A}$  will be called the beginning of a  $\tau$ -packet if it is followed by a  $\bar{\tau}$ -entrance into  $B$  from  $\bar{A}$ .

Suppose that the original flow is a superposition of  $m$  renewal flows having renewal intensities (1),  $h(t) \rightarrow h_0, t \rightarrow \infty$ . Then the subflow of initial moments of  $\varepsilon$ -packets, as  $\varepsilon \downarrow 0$ , converges in distribution to a Poisson flow in a time scale normalized by the intensity of occurrence of  $\varepsilon$ -packets. This result can be obtained by verifying conditions (8), (9) for the flow of entrances of the Markov process  $\xi_t = (u_{1t}, \dots, u_{mt})$ , where  $u_{it}$  are the backward recurrence times (1) into the set

$$A = \{(x_1, \dots, x_m), x_i = 0, x_j < \tau, x_k > 0, i \neq j \neq k, i, j, k = 1, \dots, m\}$$

from  $\bar{A}$ .

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Received  
20 XII 1968

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*Note: Figure translations are in progress. See original paper for figures.*

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