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# I. L. Romanovskaya

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**Abstract**

**Full Text**

**I. L. Romanovskaya**

## **On the Fisher–Welch–Wald Test**

*(Presented by Academician A. N. Kolmogorov on 13 I 1964)*

In the present note we consider the question of the nonexistence of the Fisher–Welch–Wald test under certain assumptions on the test boundary. Let  $x_1, x_2, \dots, x_{n_1} \in N(a_1, \sigma_1)$ ,  $y_1, y_2, \dots, y_{n_2} \in N(a_2, \sigma_2)$  be normal repeated samples with different variances. The parameters  $a_1, a_2, \sigma_1, \sigma_2$  are assumed unknown. The hypothesis  $H_0 : a_1 = a_2$  is tested by means of a nonrandomized test similar with respect to the parameter  $\theta = n_2 \sigma_1^2 / n_1 \sigma_2^2$ . The specification of the test determines a certain function  $\psi(x_1, \dots, x_{n_1}, y_1, \dots, y_{n_2})$  ( $0 \leq \psi \leq 1$ )—the probability of rejecting the hypothesis  $H_0$  after the observations. If the function  $\psi$  takes only the two values 0 or 1, then the test is called **nonrandomized**; otherwise the test is **randomized**. Tests of this kind are completely characterized by specifying a critical region  $Z$  in the Cartesian product of the sample spaces. The test will be similar if the probability of falling into the critical region  $Z$  does not depend on the parameter  $\theta$ . Using A. Wald's conditions on the critical region  $Z$ , one can show <sup>(1)</sup> that the critical region has the form:

$$\frac{|\bar{x} - \bar{y}|}{\sqrt{s_1^2 + s_2^2}} \geq \varphi\left(\frac{s_1}{s_2}\right), \quad (1)$$

where  $\bar{x}, \bar{y}, s_1^2, s_2^2$  are sufficient statistics for the four parameters  $a_1, a_2, \sigma_1^2, \sigma_2^2$ , and  $\varphi(s_1/s_2)$  is a single-valued measurable function for  $s_1/s_2 \geq 0$ . If we put  $(\bar{x} - \bar{y})/s_2 = \xi$ ,  $s_1/s_2 = \eta$  and double the corresponding probabilities of falling into the region, then we obtain the region  $Z_1$ :

$$\xi \geq \Phi(\eta),$$

where  $\xi, \eta$  vary in the right upper quadrant  $\Omega_1 : \xi \geq 0, \eta \geq 0$ , and  $\Phi(\eta) = \sqrt{1 + \eta^2} \varphi(\eta)$ . The function  $\Phi(\eta)$  has the same smoothness properties as  $\varphi(\eta)$ . Yu. V. Linnik <sup>(2)</sup>, in the case of samples of equal size and under certain smoothness conditions on the function  $\varphi(\eta)$ , showed that there is no nonrandomized similar Fisher–Welch–Wald test. In the present note the results of <sup>(2)</sup> are somewhat strengthened, and they are extended to the case of samples of arbitrary size.

Considering the joint density of the distribution of the quantities  $\xi, \eta$ , and taking into account that the test must be similar, we arrive at the basic equation of

the problem <sup>(3)</sup>:

$$\iint_{\Omega_1} \psi(\xi, \eta) \frac{\eta^{n_1-2} d\xi d\eta}{[\theta^2 + \theta(1 + \xi^2 + \eta^2) + \eta^2]^N} = C_{n_1, n_2} \frac{\alpha}{2} \theta^{-n_2/2} (1 + \theta)^{-N+1/2}, \quad (2)$$

where  $\theta = n_2 \sigma_1^2 / n_1 \sigma_2^2$ ;  $C_{n_1, n_2}$  is a constant depending on  $n_1, n_2$ ;  $\alpha/2$  is the size of the critical region  $Z_1$ ;  $N = (n_1 + n_2)/2 - 1/2$ ;

$$\psi(\xi, \eta) = \begin{cases} 1, & \text{if } (\xi, \eta) \in Z_1, \\ 0, & \text{if } (\xi, \eta) \notin Z_1. \end{cases}$$

The denominator of the fraction on the left in equation (2) vanishes only for negative values of  $\theta$ . The zeros of the denominator generate

geometric loci of points, which are called **critics** <sup>(3,4)</sup>. If in (2)  $\theta = -D_0$  is a root of the denominator, then for  $D \leq 1$  the critics  $A(\xi, \eta) = D$  form a family of confocal hyperbolas of the form  $\frac{\eta^2}{D} - \frac{\xi^2}{1-D} = 1$ . If  $D \geq 1$ , then the critics

$B(\xi, \eta) = D$  give a family of confocal semiellipses of the form  $\frac{\xi^2}{D-1} + \frac{\eta^2}{D} = 1$ .

Since both sides of (2) can be continued to the domain  $\Lambda$  of complex values of the parameter  $\theta = \tau + i\zeta$  with a cut along the axis  $\tau \leq 0$ , and the integral on the left represents there an analytic function <sup>(4)</sup>, choosing the corresponding branches of the factors for  $\theta \in \Lambda$ , we can rewrite the basic equation in the form

$$\iint_{\Omega_1} \psi(\xi, \eta) \frac{\eta^{n_1-2} d\xi d\eta}{(\theta + A(\xi, \eta))^N (\theta + B(\xi, \eta))^N} = C_{n_1, n_2} \frac{\alpha}{2} \theta^{-n_2/2} (1 + \theta)^{-N+1/2}.$$

**Theorem 1.** There does not exist a similar nonrandomized test of A. Wald of the form (1), for which  $\varphi(\eta)$  is continuous in the interval  $(0, \eta_1]$ ,  $\sup_{\eta \geq 1} \varphi(\eta) \leq c < \infty$ ;  $\varphi(\eta)$  has a finite first derivative in the interval  $[1 + \varepsilon, \eta_1]$ , where  $\varepsilon$  is a small number and  $\eta_1$  is sufficiently large, and this function satisfies a Lipschitz condition in some interval.

In the theorem obtained in <sup>(2)</sup> for the case of samples of equal size, the condition is imposed that the function  $\varphi(\eta)$  be continuous at the point  $\eta = 0$ . Here this condition is replaced by the weaker requirement that the function  $\varphi(\eta)$  be bounded as  $\eta \rightarrow +0$ .

In the proof the following lemma is used essentially.

**Lemma.** Among the critic-hyperbolas there is a critic  $A(\xi, \eta) = D_0$  ( $D_0 \in (0, 1)$ ), which is tangent to the test boundary  $\Phi(\eta) = \xi$  at some point  $(\xi_0, \eta_0)$ , with  $\xi_0 \neq 0, \eta_0 > 1$ .

The proof of the lemma is based on simple geometric considerations.

For the proof of the theorem one studies the behavior of the imaginary part of the basic equation when  $\theta = -D_0 + i\zeta$  and  $\zeta \rightarrow 0$  ( $D_0$  is a root of the denominator). The domain of integration is divided into two parts: in one of them the denominator of the left-hand side of the basic equation is separated from zero by a certain constant. In this domain the imaginary part of the integral in (2), as  $\zeta \rightarrow 0$ , is a bounded quantity; while in the domain bounded by the two hyperbolas  $A(\xi, \eta) = D_0 + \varepsilon_0$  and  $A(\xi, \eta) = D_0 - \varepsilon_0$  ( $\varepsilon_0$  sufficiently small) and containing the point of tangency of the critic  $A(\xi, \eta) = D_0$  and the test boundary, the imaginary part of the integral under consideration tends to  $-\infty$ .

Consider a randomized test in which the function  $\psi(\xi, \eta)$  represents the probability of accepting the hypothesis  $H_0$ . It is shown in (5) that randomized tests exist if the function  $\psi(\xi, \eta)$  is discontinuous.

**Theorem 2.** If the function  $\psi(\xi, \eta)$  increases monotonically in  $\xi$  for fixed  $\eta$  and is a discontinuous function, and moreover is such that the projection of the discontinuity onto the  $(\xi, \eta)$ -plane has the same smoothness properties in the interval  $[1 + \varepsilon, \eta_1]$  as the test boundary in Theorem 1, then  $\psi(\xi, \eta)$  cannot define a randomized test similar with respect to the parameter  $\theta$ .

Projecting the line of discontinuity of the function  $\psi(\xi, \eta)$  onto the  $(\xi, \eta)$ -plane leads us to the conditions of Theorem 1.

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## CITED LITERATURE

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*Note: Figure translations are in progress. See original paper for figures.*

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