



Soviet-era science, translated into English

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1957

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Abstract

Full Text

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ON SOME LIMIT THEOREMS FOR HOMOGENEOUS MARKOV CHAINS

(Presented by Academician A. N. Kolmogorov on 13 II 1957)

Let an abstract space X be given, and let \mathcal{F}_x be a σ -algebra of its subsets. Let $p(\eta, A)$, $\eta \in X$, $A \in \mathcal{F}_x$, be the transition-probability function. In what follows we shall assume that there exists a stationary probability distribution $p(A)$ such that, for some $\rho < 1$ and c ,

$$|p^{(n)}(\eta, A) - p(A)| < c\rho^n \quad (1)$$

uniformly with respect to $\eta \in X$ and $A \in \mathcal{F}_x$, where $p^{(n)}(\eta, A)$ is the probability of transition in n steps from state η to a state belonging to the set A . The function $p(\eta, A)$, together with the initial probability distribution $\pi(A)$, determines a sequence $x_1, x_2, \dots, x_n, \dots$ of random variables connected in a homogeneous Markov chain, with

$$P(x_1 \in A) = \pi(A), \quad P(x_n \in A) = \int_X p^{(n-1)}(\eta, A) \pi(d\eta). \quad (2)$$

Let $f(\eta)$ be a real function defined on X and measurable with respect to \mathcal{F}_x .

Theorem 1. If

$$\int |f(\eta)|^2 p(d\eta) < \infty,$$

$$\sigma^2 = \lim_{n \rightarrow \infty} M \left[\frac{1}{\sqrt{n}} \sum_{m=1}^n (f(x_m) - Mf(x_m)) \right] > 0$$

(the mathematical expectation is computed under the assumption that the initial distribution is stationary), then for any initial distribution $\pi(A)$

$$\lim_{n \rightarrow \infty} P \left\{ \frac{1}{\sqrt{n}} \sum_{m=1}^n \left(f(x_m) - \int_X f(\eta) p(d\eta) \right) < x \right\} = \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^x e^{-u^2/2\sigma^2} du. \quad (3)$$

This theorem is an analogue of the well-known theorem of P. Lévy, which states that if $x_1, x_2, \dots, x_n, \dots$ is a sequence of independent identically distributed random variables and $\sigma^2 = Dx_i < \infty$, then

$$\lim_{n \rightarrow \infty} P \left(\frac{\sum_{i=1}^n x_i - na}{\sigma \sqrt{n}} < x \right) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^x e^{-u^2/2} du,$$

where a is the mathematical expectation of χ_i . Up to the present time the central limit theorem for a homogeneous Markov chain with an arbitrary set of states has been proved under the assumption that, for some $\delta > 0$ (1, 2, 3),

$$\int_X |f(\eta)|^{2+\delta} p(d\eta) < \infty.$$

Theorem 2. Let $u_1, u_2, \dots, u_n, \dots$ be a sequence of independent random variables with common distribution function $F(x)$ such that

$$F(x) = p(f(\eta) < x)$$

(here $p(A)$ is the stationary distribution). If, for some sequence of constants A_n and $B_n > 0$,

$$\lim_{n \rightarrow \infty} P \left(\frac{1}{B_n} \left(\sum_{m=1}^n u_m - A_n \right) < x \right) = V_\alpha(x),$$

where $V_\alpha(x)$ is a stable law with characteristic exponent α , and if for some $0 < \nu \leq 1$

$$\lim_{n \rightarrow \infty} n^{1/2} B_n^{-\nu} \sup_{\xi} \int_{|f(\eta)| < B_n \tau} |f(\eta)|^\nu p(\xi, d\eta) = 0, \quad (4)$$

whatever $\tau > 0$ may be, then, for an arbitrary initial distribution $\pi(A)$,

$$\lim_{n \rightarrow \infty} P \left(\frac{1}{B_n} \left(\sum_{m=1}^n f(x_m) - A_n \right) < x \right) = V_\alpha(x). \quad (4')$$

Condition (4) is satisfied, for example, when for some $\varepsilon < \alpha/2$ the moments

$$\int_X |f(\eta)|^{\alpha-\varepsilon} p(\xi, d\eta)$$

are uniformly bounded in ξ .

Theorem 3. If, for some $0 < \alpha < 2$,

$$\int_X |f(\eta)|^\alpha p(d\eta) < \infty,$$

then, for some choice of the constants A_n and for an arbitrary initial distribution,

$$\lim_{n \rightarrow \infty} P \left(\frac{1}{n^{1/\alpha}} \left(\sum_{m=1}^n f(x_m) - A_n \right) < x \right) = E(x), \quad (5)$$

where $E(x)$ is an improper law.

Corollary. If

$$\int_X |f(\eta)| p(d\eta) < \infty,$$

then the sequence $f(x_1), f(x_2), \dots, f(x_n), \dots$ obeys the law of large numbers.

Let now X be a countable set $\{\omega_i\}$ ($i = 1, 2, \dots$) and

$$\beta = \inf_{(i,j)} \sum_{k=1}^{\infty} \min(p_{ik}, p_{jk}), \quad (6)$$

where p_{ik} is the probability of transition from ω_i to ω_k in one step (the meaning of condition (6) is explained in (3)).

Assume further that all states ω_i are essential and form a positive class (4). In view of (6), this class consists of a single subclass. Let $f(\omega_j) = a + k_j h$, where a is an arbitrary real number, k_j is an integer, and $h > 0$.

Theorem 4. If the greatest common divisor k_j is equal to 1,

$$\sum_{j=1}^{\infty} f^2(\omega_j) p_j < \infty$$

and $\sigma > 0$ (p_j are the final probabilities; σ is defined in the same way as in Theorem 1), then, uniformly with respect to s ,

$$\lim_{n \rightarrow \infty} \left(\frac{\sigma \sqrt{n}}{h} P_{\pi n}(s) - \frac{1}{\sqrt{2\pi}} e^{-z_{ns}^2/2} \right) = 0, \quad (7)$$

where $P_{\pi n}$ is the probability that

$$\sum_{m=1}^n f(x_m) = an + sh,$$

under the condition that the initial distribution is $\pi(A)$, and

$$z_{ns} = \frac{1}{\sigma\sqrt{n}} \left(an + sh - n \sum_{j=1}^{\infty} f(\omega_j) p_j \right).$$

Theorem 5. If the conditions of Theorem 4 are satisfied and, in addition, for some integer $k \geq 3$ and some $\delta > 0$

$$\sum_{j=1}^{\infty} |f(\omega_j)|^{k+\delta} p_{ij} < M < \infty$$

uniformly for all i , then

$$P_{\pi n}(s) = \frac{h}{\sigma\sqrt{n}} \left\{ \varphi(z_{ns}) + \sum_{m=1}^{k-2} \frac{1}{n^{m/2}} T_{\pi m}(\varphi(z_{ns})) + O\left(\frac{1}{n^{(k-2)/2}}\right) \right\}. \quad (8)$$

Here

$$\varphi(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}, \quad T_{\pi m}(\varphi(x)) = \sum_{i=m+2}^{3m} a_{\pi mi} \frac{d^i}{dx^i} \varphi(x),$$

where the coefficients $a_{\pi mi}$ depend on the initial distribution $\pi(A)$ and on $p(\eta, A)$.

Theorems 4 and 5 are a generalization of results of S. Kh. Sirazhdinov⁽⁵⁾. All the theorems formulated above are obtained by means of a method based on the application of the spectral theory of linear operators in a Banach space⁽⁷⁾. This method is a natural generalization of the matrix method and the method of integral equations, presented, for example, in⁽⁶⁾; it can also be used for proving multidimensional limit theorems.

In conclusion, the author thanks R. L. Dobrushin and V. M. Zolotarev for valuable comments.

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Received
23 XII 1957

REFERENCES

1. W. Doeblin, Bull. Math. Soc. Roum. Sci., **39**, No. 2, 3 (1937).
2. E. B. Dynkin, Ukr. Mat. Zhurn., **6**, No. 1, 21 (1954).
3. R. L. Dobrushin, *Theory of Probability and Its Applications*, **1**, 1, 72 (1956).
4. A. N. Kolmogorov, Byull. MGU, **1**, 3 (1937).
5. S. Kh. Sirazhdinov, DAN, **84**, No. 6 (1952).
6. T. A. Sarymsakov, *Foundations of the Theory of Markov Processes*, Moscow, 1954.
7. F. Riesz, B. Sz.-Nagy, *Lectures on Functional Analysis*, Moscow, 1953.
8. J. L. Doob, *Stochastic Processes*, Moscow, 1956, p. 207.

Note: Figure translations are in progress. See original paper for figures.

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